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## 11. Availability: A heuristic for judging frequency and probability

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### Introduction

Much recent research has been concerned with the validity and consistency of frequency and probability judgments. Little is known, however, about the psychological mechanisms by which people evaluate the frequency of classes or the likelihood of events.

We propose that when faced with the difficult task of judging probability or frequency, people employ a limited number of heuristics which reduce these judgments to simpler ones. Elsewhere we have analyzed in detail one such heuristic – representativeness. By this heuristic, an event is judged probable to the extent that it represents the essential features of its parent population or generating process. . . .

When judging the probability of an event by representativeness, one compares the essential features of the event to those of the structure from which it originates. In this manner, one estimates probability by assessing similarity or connotative distance. Alternatively, one may estimate probability by assessing availability, or associative distance. Life-long experience has taught us that instances of large classes are recalled better and faster than instances of less frequent classes, that likely occurrences are easier to imagine than unlikely ones, and that associative connections are strengthened when two events frequently co-occur. Thus, a person could estimate the numerosity of a class, the likelihood of an event, or the

frequency of co-occurrences by assessing the ease with which the relevant mental operation of retrieval, construction, or association can be carried out.

For example, one may assess the divorce rate in a given community by recalling divorces among one's acquaintances; one may evaluate the probability that a politician will lose an election by considering various ways in which he may lose support; and one may estimate the probability that a violent person will "see" beasts of prey in a Rorschach card by assessing the strength of association between violence and beasts of prey. In all these cases, the estimation of the frequency of a class or the probability of an event is mediated by an assessment of availability.<sup>1</sup> A person is said to employ the availability heuristic whenever he estimates frequency or probability by the ease with which instances or associations could be brought to mind. To assess availability it is not necessary to perform the actual operations of retrieval or construction. It suffices to assess the ease with which these operations could be performed, much as the difficulty of a puzzle or mathematical problem can be assessed without considering specific solutions.

That associative bonds are strengthened by repetition is perhaps the oldest law of memory known to man. The availability heuristic exploits the inverse form of this law, that is, it uses strength of association as a basis for the judgment of frequency. In this theory, availability is a mediating variable, rather than a dependent variable as is typically the case in the study of memory. Availability is an ecologically valid clue for the judgment of frequency because, in general, frequent events are easier to recall or imagine than infrequent ones. However, availability is also affected by various factors which are unrelated to actual frequency. If the availability heuristic is applied, then such factors will affect the perceived frequency of classes and the subjective probability of events. Consequently, the use of the availability heuristic leads to systematic biases.

This paper explores the availability heuristic in a series of ten studies.<sup>2</sup> We first demonstrate that people can assess availability with reasonable speed and accuracy. Next, we show that the judged frequency of classes is biased by the availability of their instances for construction, and retrieval. The experimental studies of this paper are concerned with judgments of frequencies, or of probabilities that can be readily reduced to relative

<sup>1</sup> The present use of the term "availability" does not coincide with some usages of this term in the verbal learning literature (see, e.g., Horowitz, Norman, & Day, 1966; Tulving & Pearlstone, 1966).

<sup>2</sup> Approximately 1500 subjects participated in these studies. Unless otherwise specified, the studies were conducted in groups of 20-40 subjects. Subjects in Studies 1, 2, 3, 9 and 10 were recruited by advertisements in the student newspaper at the University of Oregon. Subjects in Study 8 were similarly recruited at Stanford University. Subjects in Studies 5, 6 and 7 were students in the 10th and 11th grades of several college-preparatory high schools in Israel.

frequencies. The effects of availability on the judged probabilities of essentially unique events (which cannot be reduced to relative frequencies) are discussed in the fifth and final section.

### Assessments of availability

#### *Study 1: Construction*

The subjects ( $N = 42$ ) were presented with a series of word-construction problems. Each problem consisted of a  $3 \times 3$  matrix containing nine letters from which words of three letters or more were to be constructed. In the training phase of the study, six problems were presented to all subjects. For each problem, they were given 7 sec to estimate the number of words which they believed they could produce in 2 min. Following each estimate, they were given two minutes to write down (on numbered lines) as many words as they could construct from the letters in the matrix. Data from the training phase were discarded. In the test phase, the construction and estimation tasks were separated. Each subject estimated for eight problems the number of words which he believed he could produce in 2 min. For eight other problems, he constructed words without prior estimation. Estimation and construction problems were alternated. Two parallel booklets were used, so that for each problem half the subjects estimated and half the subjects constructed words.

*Results.* The mean number of words produced varied from 1.3 (for XUZONLCJM) to 22.4 (for TAPCERHOB), with a grand mean of 11.9. The mean number estimated varied from 4.9 to 16.0 (for the same two problems), with a grand mean of 10.3. The product-moment correlation between estimation and production, over the sixteen problems, was 0.96.

#### *Study 2: Retrieval*

The design and procedure were identical to Study 1, except for the nature of the task. Here, each problem consisted of a category, e.g., *flowers* or *Russian novelists*, whose instances were to be recalled. The subjects ( $N = 28$ ) were given 7 sec to estimate the number of instances they could retrieve in 2 min, or 2 min to actually retrieve the instances. As in Study 1, the production and estimation tasks were combined in the training phase and alternated in the test phase.

*Results.* The mean number of instances produced varied from 4.1 (city names beginning with *F*) to 23.7 (four-legged animals), with a grand mean of 11.7. The mean number estimated varied from 6.7 to 18.7 (for the same two categories), with a grand mean of 10.8. The product-moment correlation between production and estimation over the 16 categories was 0.93.

*Discussion*

In the above studies, the availability of instances could be measured by the total number of instances retrieved or constructed in any given problem.<sup>3</sup> The studies show that people can assess availability quickly and accurately. How are such assessments carried out? One plausible mechanism is suggested by the work of Bousfield and Sedgewick (1944), who showed that cumulative retrieval of instances is a negatively accelerated exponential function of time. The subject could, therefore, use the number of instances retrieved in a short period to estimate the number of instances that could be retrieved in a much longer period of time. Alternatively, the subject may assess availability without explicitly retrieving or constructing any instances at all. Hart (1967), for example, has shown that people can accurately assess their ability to recognize items that they cannot recall in a test of paired-associate memory.

**Availability for construction**

We turn now to a series of problems in which the subject is given a rule for the construction of instances and is asked to estimate their total (or relative) frequency. In these problems – as in most estimation problems – the subject cannot construct and enumerate all instances. Instead, we propose, he attempts to construct some instances and judges overall frequency by availability, that is, by an assessment of the ease with which instances could be brought to mind. As a consequence, classes whose instances are easy to construct or imagine will be perceived as more frequent than classes of the same size whose instances are less available. This prediction is tested in the judgment of word frequency, and in the estimation of several combinatorial expressions.

*Study 3: Judgment of word frequency*

Suppose you sample a word at random from an English text. Is it more likely that the word starts with a *K*, or that *K* is its third letter? According to our thesis, people answer such a question by comparing the availability of the two categories, i.e., by assessing the ease with which instances of the two categories come to mind. It is certainly easier to think of words that start with a *K* than of words where *K* is in the third position. If the judgment of frequency is mediated by assessed availability, then words

<sup>3</sup> Word-construction problems can also be viewed as retrieval problems because the response-words are stored in memory. In the present paper we speak of retrieval when the subject recalls instances from a natural category, as in Studies 2 and 8. We speak of construction when the subject generates exemplars according to a specified rule, as in Studies 1 and 4.

that start with *K* should be judged more frequent. In fact, a typical text contains twice as many words in which *K* is in the third position than words that start with *K*.

According to the extensive word-count of Mayzner and Tresselt (1965), there are altogether eight consonants that appear more frequently in the third than in the first position. Of these, two consonants (*X* and *Z*) are relatively rare, and another (*D*) is more frequent in the third position only in three-letter words. The remaining five consonants (*K, L, N, R, V*) were selected for investigation.

The subjects were given the following instructions:

The frequency of appearance of letters in the English language was studied. A typical text was selected, and the relative frequency with which various letters of the alphabet appeared in the first and third positions in words was recorded. Words of less than three letters were excluded from the count.

You will be given several letters of the alphabet, and you will be asked to judge whether these letters appear more often in the first or in the third position, and to estimate the ratio of the frequency with which they appear in these positions.

A typical problem read as follows:

Consider the letter *R*.

Is *R* more likely to appear in

— the first position?

— the third position? (check one)

My estimate for the ratio of these two values is \_\_\_\_: 1.

Subjects were instructed to estimate the ratio of the larger to the smaller class. For half the subjects, the ordering of the two positions in the question was reversed. In addition, three different orderings of the five letters were employed.

*Results.* Among the 152 subjects, 105 judged the first position to be more likely for a majority of the letters, and 47 judged the third position to be more likely for a majority of the letters. The bias favoring the first position is highly significant ( $p < .001$ , by sign test). Moreover, each of the five letters was judged by a majority of subjects to be more frequent in the first than in the third position. The median estimated ratio was 2:1 for each of the five letters. These results were obtained despite the fact that all letters were more frequent in the third position.

In other studies we found the same bias favoring the first position in a within-subject design where each subject judged a single letter, and in a between-subjects design, where the frequencies of letters in the first and in the third positions were evaluated by different subjects. We also observed that the introduction of payoffs for accuracy in the within-subject design had no effect whatsoever. Since the same general pattern of

results was obtained in all these methods, only the findings obtained by the simplest procedure are reported here.

A similar result was reported by Phillips (1966) in a study of Bayesian inference. Six editors of a student publication estimated the probabilities that various bigrams, sampled from their own writings, were drawn from the beginning or from the end of words. An incidental effect observed in that study was that all the editors shared a common bias to favor the hypothesis that the bigrams had been drawn from the beginning of words. For example, the editors erroneously judged words beginning with *re* to be more frequent than words ending with *re*. The former, of course, are more available than the latter.

#### *Study 4: Permutations*

Consider the two structures, A and B, which are displayed below.

(A)	(B)
x x x x x x x	x x
x x x x x x x	x x
x x x x x x x	x x
	x x
	x x
	x x
	x x
	x x
	x x

A path in a structure is a line that connects an element in the top row to an element in the bottom row, and passes through one and only one element in each row.

In which of the two structures are there more paths?

How many paths do you think there are in each structure?

Most readers will probably share with us the immediate impression that there are more paths in A than in B. Our subjects agreed: 46 of 54 respondents saw more paths in A than in B ( $p < .001$ , by sign test). The median estimates were 40 paths in A and 18 in B. In fact, the number of paths is the same in both structures, for  $8^3 = 2^9 = 512$ .

Why do people see more paths in A than in B? We suggest that this result reflects the differential availability of paths in the two structures. There are several factors that make the paths in A more available than those in B. First, the most immediately available paths are the columns of the structures. There are 8 columns in A and only 2 in B. Second, among the paths that cross columns, those of A are generally more distinctive and less confusable than those in B. Two paths in A share, on the average, about 1/8 of their elements, whereas two paths in B share, on the average, half of their elements. Finally, the paths in A are shorter and hence easier to visualize than those in B.

### Study 5: Combinations

Consider a group of ten people who have to form committees of  $r$  members, where  $r$  is some number between 2 and 8. How many different committees of  $r$  members can they form? The correct answer to this problem is given by the binomial coefficient  $\binom{10}{r}$ , which reaches a maximum of 252 for  $r = 5$ . Clearly, the number of committees of  $r$  members equals the number of committees of  $10 - r$  members because any elected group of, say, two members defines a unique nonelected group of eight members.

According to our analysis of intuitive estimation, however, committees of two members are more available than committees of eight. First, the simplest scheme for constructing committees is a partition of the group into disjoint subsets. Thus, one readily sees that there are as many as five disjoint committees of two members, but not even two disjoint committees of eight. Second, committees of eight members are much less distinct, because of their overlapping membership; any two committees of eight share at least six members. This analysis suggests that small committees are more available than large committees. By the availability hypothesis, therefore, the small committees should appear more numerous.

Four groups of subjects (total  $N = 118$ ) estimated the number of possible committees of  $r$  members that can be formed from a set of ten people. The different groups, respectively, evaluated the following values of  $r$ : 2 and 6; 3 and 8; 4 and 7; 5.

Median estimates of the number of committees are shown in Figure 1, with the correct values. As predicted, the judged numerosity of committees decreases with their size.

The following alternative formulation of the same problem was devised in order to test the generality of the findings:

In the drawing below, there are ten stations along a route between Start and Finish. Consider a bus that travels, stopping at exactly  $r$  stations along this route.

START 

--	--	--	--	--	--	--	--	--	--

 FINISH

What is the number of different patterns of  $r$  stops that the bus can make?

The number of different patterns of  $r$  stops is again given by  $\binom{10}{r}$ . Here too, of course, the number of patterns of two stops is the same as the number of patterns of eight stops, because for any pattern of stops there is a unique complementary pattern of non-stops. Yet, it appears as though one has more degrees of freedom in constructing patterns of two stops where "one has many stations to choose from" than in constructing patterns of eight stops where "one must stop at almost every station." Our previous analysis suggests that the former patterns are more available: more such patterns are seen at first glance, they are more distinctive, and they are easier to visualize.

Four new groups of subjects (total  $N = 178$ ) answered this question, for

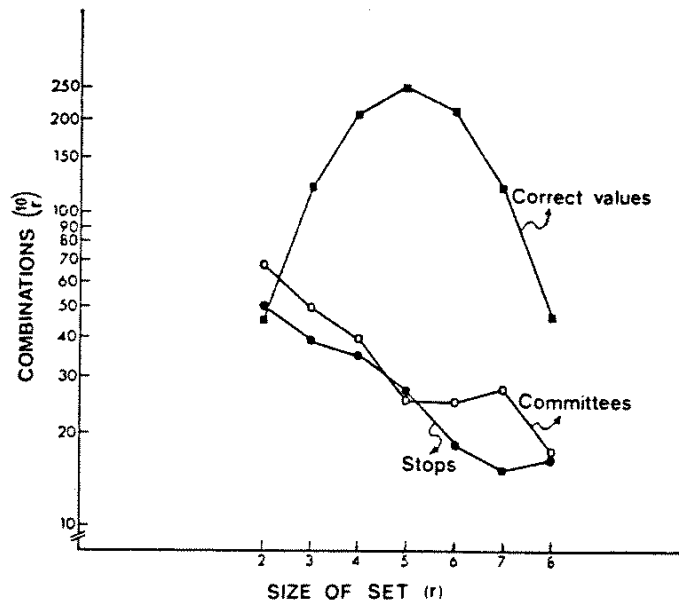


Figure 1. Correct values and median judgments (on a logarithmic scale) for the Committees problem and for the Stops problem.

$r = 2, \dots, 8$ , following the same design as above. Median estimates of the number of stops are shown in Figure 1. As in the committee problem, the apparent number of combinations generally decreases with  $r$ , in accordance with the prediction from the availability hypothesis, and in marked contrast to the correct values. Further, the estimates of the number of combinations are very similar in the two problems. As in other combinatorial problems, there is marked underestimation of all correct values, with a single exception in the most available case, where  $r = 2$ .

The underestimation observed in Experiments 4 and 5 occurs, we suggest, because people estimate combinatorial values by extrapolating from an initial impression. What a person sees at a glance or in a few steps of computation gives him an inadequate idea of the explosive rate of growth of many combinatorial expressions. In such situations, extrapolating from an initial impression leads to pronounced underestimation. This is the case whether the basis for extrapolation is the initial availability of instances, as in the preceding two studies, or the output of an initial computation, as in the following study.

#### *Study 6: Extrapolation*

We asked subjects to estimate, within 5 sec, a numerical expression that was written on the blackboard. One group of subjects ( $N = 87$ ) estimated the product  $8 \times 7 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1$ , while another group ( $N = 114$ ) estimated the product  $1 \times 2 \times 3 \times 4 \times 5 \times 6 \times 7 \times 8$ . The median estimate



or the descending sequence was 2,250. The median estimate for the ascending sequence was 512. The difference between the estimates is highly significant ( $p < .001$ , by median test). Both estimates fall very short of the correct answer, which is 40,320.

Both the underestimation of the correct value and the difference between the two estimates support the hypothesis that people estimate  $8!$  by extrapolating from a partial computation. The factorial, like other combinatorial expressions, is characterized by an ever-increasing rate of growth. Consequently, a person who extrapolates from a partial computation will grossly underestimate factorials. Because the results of the first few steps of multiplication (performed from left to right) are larger in the ascending sequence than in the descending sequence, the former expression is judged larger than the latter. The evaluation of the descending sequence may proceed as follows: "8 times 7 is 56 times 6 is already above 100, so we are dealing with a reasonably large number." In evaluating the ascending sequence, on the other hand, one may reason: "1 times 2 is 2 times 3 is 6 times 4 is 24, and this expression is clearly not going very far. . . ."

#### *Study 7: Binomial – availability vs. representativeness*

The final study of this section explores the role of availability in the valuation of binomial distributions and illustrates how the formulation of a problem controls the choice of the heuristic that people adopt in intuitive estimation.

The subjects ( $N = 73$ ) were presented with these instructions:

Consider the following diagram:

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X X O X X X
X X X X O X
X O X X X X
X X X O X X
X X X X X O
O X X X X X

```

A path in this diagram is any descending line which starts at the top row, ends at the bottom row, and passes through exactly one symbol (X or O) in each row.

What do you think is the percentage of paths which contain

6 - X and no - O \_\_\_\_\_%

5 - X and 1 - O \_\_\_\_\_%

.

.

.

No - X and 6 - O \_\_\_\_\_%

Note that these include all possible path-types and hence your estimates should add to 100%.

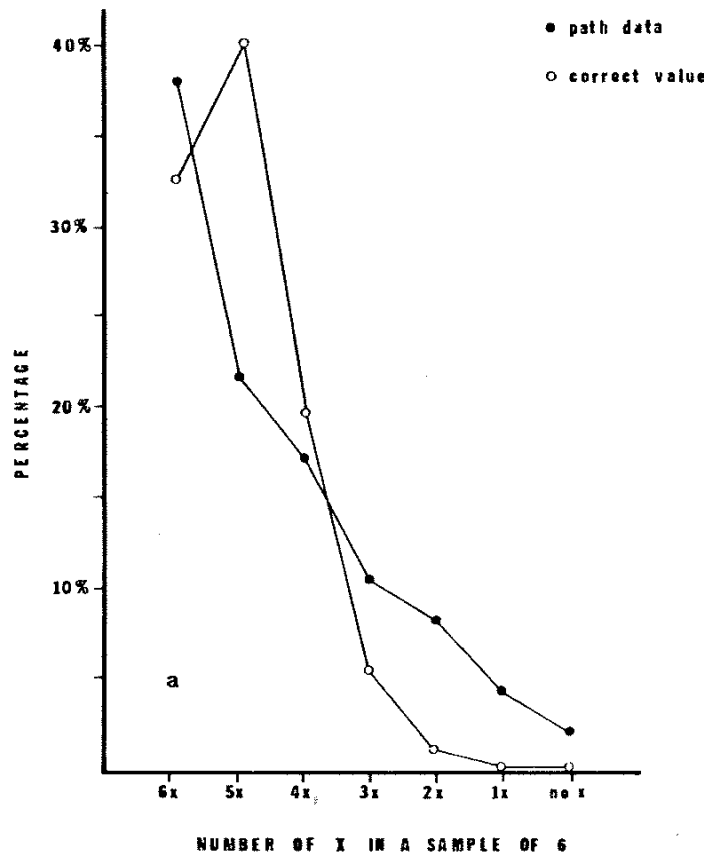


Figure 2. Correct values and median judgments: Path problem.

The actual distribution of path-type is binomial with  $p = 5/6$  and  $n = 6$ . People, of course, can neither intuit the correct answers nor enumerate all relevant instances. Instead, we propose, they glance at the diagram and estimate the relative frequency of each path-type by the ease with which individual paths of this type could be constructed. Since, at every stage in the construction of a path (i.e., in each row of the diagram) there are many more Xs than Os, it is easier to construct paths consisting of six Xs than paths consisting of, say, five Xs and one O, although the latter are, in fact, more numerous. Accordingly, we predicted that subjects would erroneously judge paths of 6 Xs and no O to be the most numerous.

Median estimates of the relative frequency of all path-types are presented in Figure 2, along with the correct binomial values. The results confirm the hypothesis. Of the 73 subjects, 54 erroneously judged that there are more paths consisting of six Xs and no O than paths consisting of five Xs and one O, and only 13 regarded the latter as more numerous than the former ( $p < .001$ , by sign test). The monotonicity of the subjective distribution of path-types is apparently a general phenomenon. We have obtained the same result with different values of  $p$  ( $4/5$  and  $5/6$ ) and  $n$  (5, 6 and 10), and different representations of the population proportions (e.g.,

four Xs and one O or eight Xs and two Os in each row of the path diagram).

To investigate further the robustness of this effect, the following additional test was conducted. Fifty combinatorially naive undergraduates from Stanford University were presented with the path problem. Here, the subjects were not asked to estimate relative frequency but merely to judge "whether there are more paths containing six Xs and no O, or more paths containing five Xs and one O." The subjects were run individually, and they were promised a \$1 bonus for a correct judgment. The significant majority of subjects (38 of 50,  $p < .001$ , by sign test) again selected the former outcome as more frequent. Erroneous intuitions, apparently, are not easily rectified by the introduction of monetary payoffs.

We have proposed that when the binomial distribution is represented as a path diagram, people judge the relative frequency of the various outcomes by assessing the availability of individual paths of each type. This mode of evaluation is suggested by the sequential character of the definition of a path and by the pictorial representation of the problem. Consider next an alternative formulation of the same problem.

Six players participate in a card game. On each round of the game, each player receives a single card drawn blindly from a well-shuffled deck. In the deck, 5/6 of the cards are marked X and the remaining 1/6 are marked O. In many rounds of the game, what is the percentage of rounds in which

6 players receive X and no player receives O    \_\_\_%

5 players receive X and 1 player receives O    \_\_\_%

.

.

.

No player receives X and 6 players receive O    \_\_\_%

Note that these include all the possible outcomes and hence your estimates should add to 100%.

This card problem is formally identical to the path problem, but it is intended to elicit a different mode of evaluation. In the path problem, individual instances were emphasized by the display, and the population proportion (i.e., the proportion of Xs in each row) was not made explicit. In the card problem, on the other hand, the population proportion is explicitly stated and no mention is made of individual instances. Consequently, we hypothesize that the outcomes in the card problem will be evaluated by the degree to which they are representative of the composition of the deck rather than by the availability of individual instances. In the card problem, the outcome "five Xs and one O" is the most representative, because it matches the population proportion (see Kahneman & Tversky, 1972b, 3). Hence, by the representativeness heuristic, this outcome should be judged more frequent than the outcome "six Xs and no O," contrary to the observed pattern of judgments in the path problem.

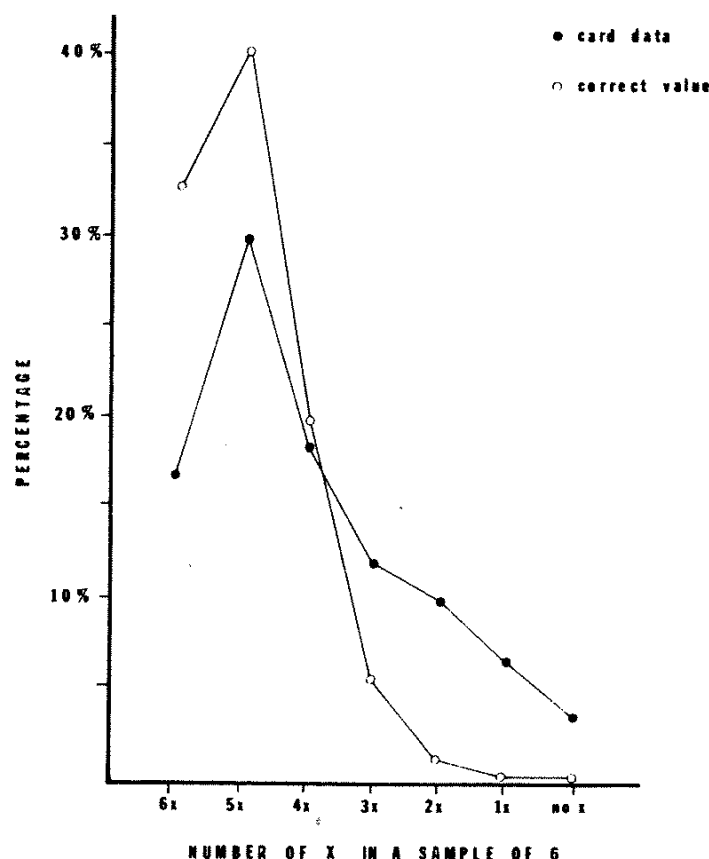


Figure 3. Correct values and median judgments: Card problem.

The judgments of 71 of 82 subjects who answered the card problem conformed to this prediction. In the path problem, only 13 of 73 subjects had judged these outcomes in the same way; the difference between the two versions is highly significant ( $p < .001$ , by a  $\chi^2$  test).

Median estimates for the card problem are presented in Figure 3. The contrast between Figures 2 and 3 supports the hypothesis that different representations of the same problem elicit different heuristics. Specifically, the frequency of a class is likely to be judged by availability if the individual instances are emphasized and by representativeness if generic features are made salient.

### Availability for retrieval

In this section we discuss several studies in which the subject is first exposed to a message (e.g., a list of names) and is later asked to judge the frequency of items of a given type that were included in the message. As in the problems studied in the previous section, the subject cannot recall and count all instances. Instead, we propose, he attempts to recall some instances and judges overall frequency by availability, i.e., by the ease

with which instances come to mind. As a consequence, classes whose instances are readily recalled will be judged more numerous than classes of the same size whose instances are less available. This prediction is first tested in a study of the judged frequency of categories. . . .

*Study 8: Fame, frequency, and recall*

The subjects were presented with a recorded list consisting of names of known personalities of both sexes. After listening to the list, some subjects judged whether it contained more names of men or of women, others attempted to recall the names in the list. Some of the names in the list were very famous (e.g., Richard Nixon, Elizabeth Taylor), others were less famous (e.g., William Fulbright, Lana Turner). Famous names are generally easier to recall. Hence, if frequency judgments are mediated by assessed availability, then a class consisting of famous names should be judged more numerous than a comparable class consisting of less famous names.

Four lists of names were prepared, two lists of entertainers and two lists of other public figures. Each list included 39 names recorded at a rate of one name every 2 sec. Two of the lists (one of public figures and one of entertainers) included 19 names of famous women and 20 names of less famous men. The two other lists consisted of 19 names of famous men and 20 names of less famous women. Hence, fame and frequency were inversely related in all lists. The first names of all personalities always permitted an unambiguous identification of sex.

The subjects were instructed to listen attentively to a recorded message. Each of the four lists was presented to two groups. After listening to the recording, subjects in one group were asked to write down as many names as they could recall from the list. The subjects in the other group were asked to judge whether the list contained more names of men or of women.

*Results.* (a) Recall. On the average, subjects recalled 12.3 of the 19 famous names and 8.4 of the 20 less famous names. Of the 86 subjects in the four recall groups, 57 recalled more famous than nonfamous names, and only 13 recalled fewer famous than less famous names ( $p < .001$ , by sign test).

(b) Frequency. Among the 99 subjects who compared the frequency of men and women in the lists, 80 erroneously judged the class consisting of the more famous names to be more frequent ( $p < .001$ , by sign test). . . .

### **Retrieval of occurrences and construction of scenarios**

In all the empirical studies that were discussed in this paper, there existed an objective procedure for enumerating instances (e.g., words that begin with *K* or paths in a diagram), and hence each of the problems had an

objectively correct answer. This is not the case in many real-life situations where probabilities are judged. Each occurrence of an economic recession, a successful medical operation, or a divorce, is essentially unique, and its probability cannot be evaluated by a simple tally of instances. Nevertheless, the availability heuristic may be applied to evaluate the likelihood of such events.

In judging the likelihood that a particular couple will be divorced, for example, one may scan one's memory for similar couples which this question brings to mind. Divorce will appear probable if divorces are prevalent among the instances that are retrieved in this manner. Alternatively, one may evaluate likelihood by attempting to construct stories, or scenarios, that lead to a divorce. The plausibility of such scenarios, or the ease with which they come to mind, can provide a basis for the judgment of likelihood. In the present section, we discuss the role of availability in such judgments, speculate about expected sources of bias, and sketch some directions that further inquiry might follow.

We illustrate availability biases by considering an imaginary clinical situation.<sup>4</sup> A clinician who has heard a patient complain that he is tired of life, and wonders whether that patient is likely to commit suicide may well recall similar patients he has known. Sometimes only one relevant instance comes to mind, perhaps because it is most memorable. Here, subjective probability may depend primarily on the similarity between that instance and the case under consideration. If the two are very similar, then one expects that what has happened in the past will recur. When several instances come to mind, they are probably weighted by the degree to which they are similar, in essential features, to the problem at hand.

How are relevant instances selected? In scanning his past experience does the clinician recall patients who resemble the present case, patients who attempted suicide, or patients who resemble the present case *and* attempted suicide? From an actuarial point of view, of course, the relevant class is that of patients who are similar, in some respects, to the present case, and the relevant statistic is the frequency of attempted suicide in this class.

Memory search may follow other rules. Since attempted suicide is a dramatic and salient event, suicidal patients are likely to be more memorable and easier to recall than depressive patients who did not attempt suicide. As a consequence, the clinician may recall suicidal patients he has encountered and judge the likelihood of an attempted suicide by the degree of resemblance between these cases and the present patient. This approach leads to serious biases. The clinician who notes that nearly all suicidal patients he can think of were severely depressed may conclude

<sup>4</sup> This example was chosen because of its availability. We know of no reason to believe that intuitive predictions of stockbrokers, sportscasters, political analysts or research psychologists are less susceptible to biases.

that a patient is likely to commit suicide if he shows signs of severe depression. Alternatively, the clinician may conclude that suicide is unlikely if "this patient does not look like any suicide case I have met." Such reasoning ignores the fact that only a minority of depressed patients attempt suicide and the possibility that the present patient may be quite unlike any that the therapist has ever encountered.

Finally, a clinician might think only of patients who were both depressed and suicidal. He would then evaluate the likelihood of suicide by the ease with which such cases come to mind or by the degree to which the present patient is representative of this class. This reasoning, too, is subject to a serious flaw. The fact that there are many depressed patients who attempted suicide does not say much about the probability that a depressed patient will attempt suicide, yet this mode of evaluation is not uncommon. Several studies (Jenkins & Ward, 1963; Smedslund, 1963; Ward & Jenkins, 1965) showed that contingency between two binary variables such as a symptom and a disease is judged by the frequency with which they co-occur, with little or no regard for cases where either the symptom or the disease was not present.

Some events are perceived as so unique that past history does not seem relevant to the evaluation of their likelihood. In thinking of such events we often construct *scenarios*, i.e., stories that lead from the present situation to the target event. The plausibility of the scenarios that come to mind, or the difficulty of producing them, then serve as a clue to the likelihood of the event. If no reasonable scenario comes to mind, the event is deemed impossible or highly unlikely. If many scenarios come to mind, or if the one scenario that is constructed is particularly compelling, the event in question appears probable.

Many of the events whose likelihood people wish to evaluate depend on several interrelated factors. Yet it is exceedingly difficult for the human mind to apprehend sequences of variations of several interacting factors. We suggest that in evaluating the probability of complex events only the simplest and most available scenarios are likely to be considered. In particular, people will tend to produce scenarios in which many factors do not vary at all, only the most obvious variations take place, and interacting changes are rare. Because of the simplified nature of imagined scenarios, the outcomes of computer simulations of interacting processes are often counter-intuitive (Forrester, 1971). The tendency to consider only relatively simple scenarios may have particularly salient effects in situations of conflict. There, one's own moods and plans are more available to one than those of the opponent. It is not easy to adopt the opponent's view of the chessboard or of the battlefield, which may be why the mediocre player discovers so many new possibilities when he switches sides in a game. Consequently, the player may tend to regard his opponent's strategy as relatively constant and independent of his own moves. These considerations suggest that a player is susceptible to the *fallacy of initiative*

- a tendency to attribute less initiative and less imagination to the opponent than to himself. This hypothesis is consistent with a finding of attribution-research (Jones & Nisbett, 1971) that people tend to view their own behavior as reflecting the changing demands of their environment and others' behavior as trait-dominated.

The production of a compelling scenario is likely to constrain future thinking. There is much evidence showing that, once an uncertain situation has been perceived or interpreted in a particular fashion, it is quite difficult to view it in any other way (see, e.g., Bruner & Potter, 1969). Thus, the generation of a specific scenario may inhibit the emergence of other scenarios, particularly those that lead to different outcomes. . . .

Perhaps the most obvious demonstration of availability in real life is the impact of the fortuitous availability of incidents or scenarios. Many readers must have experienced the temporary rise in the subjective probability of an accident after seeing a car overturned by the side of the road. Similarly, many must have noticed an increase in the subjective probability that an accident or malfunction will start a thermonuclear war after seeing a movie in which such an occurrence was vividly portrayed. Continued preoccupation with an outcome may increase its availability, and hence its perceived likelihood. People are preoccupied with highly desirable outcomes, such as winning the sweepstakes, or with highly undesirable outcomes, such as an airplane crash. Consequently, availability provides a mechanism by which occurrences of extreme utility (or disutility) may appear more likely than they actually are. . . .